

# Advanced Econometrics I: Principles of Econometrics

## Final Exam

16 December 2016

This exam has three pages. Good luck!

### 1 Gauß-Markov Theorem (12 points)

Prove the Gauß-Markov Theorem: The OLS estimator  $b = (X'X)^{-1} X'y$  is the BLUE for  $\beta$  in the linear model  $y = X\beta + \varepsilon$ , and the variance of  $b$  is  $\sigma^2 (X'X)^{-1}$ .

#### 1.1 Statement and definitions (4 points)

1. State the Gauß-Markov assumptions (you may assume that the matrix  $X \in \mathbb{R}^{N \times K}$ ,  $N \geq K$ , is non-stochastic).
2. Explain each letter in the acronym "BLUE" and give definitions of the terms involved.

#### 1.2 Part 1 of the proof (4 points)

1. Derive the conditions on the matrix  $D$  in the linear estimator  $\tilde{b} = Dy$  such that it satisfies property "U" in "BLUE".
  - (a) What dimensions does  $D$  have?
2. Verify that the OLS estimator  $b$  satisfies the derived condition.

#### 1.3 Part 2 of the proof (4 points)

1. Show that for matrices  $D$ ,  $X \in \mathbb{R}^{N \times K}$ ,  $I$  (the identity matrix) satisfying  $DX = I$ , the decomposition

$$DD' = \left( (X'X)^{-1} X' \right) \left( (X'X)^{-1} X' \right)' + \left( D - \left[ (X'X)^{-1} X' \right] \right) \left( D - \left[ (X'X)^{-1} X' \right] \right)'$$

holds.

2. Show that the OLS estimator satisfies "B" in the acronym BLUE.

## 2 F-test (12 points)

Consider the regression model

$$y = X_1\beta^{(1)} + X_2\beta^{(2)} + \varepsilon$$

satisfying assumptions (A1)-(A5). Furthermore, assume that  $X_1 \in \mathbb{R}^{N \times (K-J)}$  and  $X_2 \in \mathbb{R}^{N \times J}$  are non-stochastic.

Derive the F-test for  $H_0 : \beta_{K-J+1} = \dots = \beta_K = 0$ , i.e. show that

$$\frac{\frac{1}{\sigma^2} (e'_{restr} e_{restr} - e'e)}{\frac{1}{\sigma^2} e'e} \sim F_{J, N-K}$$

where  $e$  is the OLS residual in the unrestricted model, and  $e_{restr}$  is the OLS residual under  $H_0$ , i.e. in the restricted model.

1. (3 points) Assuming  $\frac{1}{\sigma^2} (e'_{restr} e_{restr} - e'e)$  and  $\frac{1}{\sigma^2} e'e$  are independent, why is  $f = \frac{(e'_{restr} e_{restr} - e'e)}{\frac{1}{N-K} e'e}$  distributed as  $F_{J, N-K}$ ?
2. (3 points) Project the equation  $y = X_1 b^{(1)} + X_2 b^{(2)} + e$ , where  $b = \begin{pmatrix} b^{(1)} \\ b^{(2)} \end{pmatrix}$  is the OLS estimator for  $\beta = \begin{pmatrix} \beta^{(1)} \\ \beta^{(2)} \end{pmatrix}$ , on the orthogonal complement of the space spanned by the columns of  $X_1$  and show that  $e'_{restr} e_{restr} = (b^{(2)})' \tilde{X}'_2 \tilde{X}_2 b^{(2)} + e'e$ . Explain your calculations.
3. (3 points) Show that under  $H_0$  the equality  $\frac{1}{\sigma^2} (e'_{restr} e_{restr} - e'e) = \frac{1}{\sigma^2} [\varepsilon' P_{\tilde{X}_2} \varepsilon]$  holds. (You may use the Frisch-Waugh Theorem, i.e. use the fact that  $b^{(2)} = \left( \begin{pmatrix} \tilde{X}_2 \\ \tilde{X}_2 \end{pmatrix}' \tilde{X}_2 \right)^{-1} \begin{pmatrix} \tilde{X}_2 \end{pmatrix}' y$ .) Explain your calculations.
4. (3 points) Show that  $\frac{1}{\sigma^2} e'e$  is equal to  $\frac{1}{\sigma^2} \varepsilon' M_X \varepsilon$ . Then, show that  $\frac{1}{\sigma^2} \varepsilon' M_X \varepsilon$  and  $\frac{1}{\sigma^2} [\varepsilon' P_{\tilde{X}_2} \varepsilon]$  are independent.
  - (a) Hint: In order to show that  $\frac{1}{\sigma^2} \varepsilon' M_X \varepsilon$  and  $\frac{1}{\sigma^2} [\varepsilon' P_{\tilde{X}_2} \varepsilon]$  are independent, start by showing that  $M_X \varepsilon$  and  $P_X \varepsilon$  are uncorrelated, then use properties of normal distributions to show that the quadratic forms above are independent.

### 3 Short Questions (12 points + 2 points BONUS)

We consider the linear regression model

$$y = X\beta + \varepsilon$$

where  $X \in \mathbb{R}^{N \times K}$  is non-stochastic,  $\mathbb{E}(\varepsilon) = 0$  and  $\mathbb{V}(\varepsilon) = \sigma^2 \Psi$ , where  $\sigma^2 > 0$  and  $\Psi$  is positive definite. Furthermore,  $\frac{1}{N} (X'X) \xrightarrow{N \rightarrow \infty} \Sigma_{xx}$  is positive-definite.

- (3 points) Calculate the covariance matrix  $\mathbb{V}(b)$  of the OLS estimator  $b = (X'X)^{-1} X'y$ .
- (1 point) The OLS estimator is unbiased in this model. True or false?
- (1 point) The OLS estimator is consistent in this model. True or false?

Consider, under the assumptions above, the case of heteroskedasticity, i.e.  $\mathbb{V}(\varepsilon) = \begin{pmatrix} \sigma_1^2 & 0 & \dots & 0 \\ 0 & \sigma_2^2 & \dots & \vdots \\ \vdots & \dots & \dots & 0 \\ 0 & \dots & 0 & \sigma_N^2 \end{pmatrix}$ .

- (1 point) The weighted least squares (WLS) estimator gives a different weight to each column of  $X$ . True or false?
- (1 point) The matrix  $\left(\frac{1}{N} \sum_{i=1}^N e_i^2 x_i x_i'\right)$  in  $\hat{\mathbb{V}}(b) = \frac{1}{N} \left(\frac{1}{N} X'X\right)^{-1} \left(\frac{1}{N} \sum_{i=1}^N e_i^2 x_i x_i'\right) \left(\frac{1}{N} X'X\right)^{-1}$  is a consistent estimator for  $\left(\frac{1}{N} \sum_{i=1}^N \sigma_i^2 x_i x_i'\right)$ . True or false?
- (1 point) How many free parameters are in  $\left(\frac{1}{N} \sum_{i=1}^N \sigma_i^2 x_i x_i'\right)$ ? Why is this advantageous compared to estimating  $\sigma_i^2$  for all  $i$  (give a numerical example with  $K = 5$  and  $N = 200$ )?

Consider the  $H_0 : R\beta = q$ . Let  $\hat{\theta}$  be the unrestricted MLE, and let  $\bar{\theta}$  be the restricted MLE.

- (1 point) The unrestricted MLE  $\hat{\theta}$  is unbiased, asymptotically efficient, and asymptotically normal. True or false?
- (1 point) Define the Wald test statistic and give a geometric interpretation. Which model needs to be estimated?
- (1 point) Define the LR test statistic and give a geometric interpretation. Which model needs to be estimated?
- (2 points) Define the LM test statistic and give a geometric interpretation. Which model needs to be estimated?
- (1 point) The LR test is sensitive with respect to the formulation of the restriction. True or false?